can require a number of spin flips that is exponential in N (A. Haken et al. ca. 1989), and that one can in fact embed arbitrary computations in the dynamics (Orponen 1995). (More precisely, determining whether a given "output spin" is +1 or -1 in the local minimum reached from a given initial state is a "PSPACE-complete" problem.)

6.4 The NK Model

Introduced by Stuart Kauffman (ca. 1986) as a "tunable family of fitness land-scapes".

A fitness landscape is a triple $\langle X, R, f \rangle$, where X is the configuration (or state) space, $R \subseteq X \times X$ is a neighbourhood relation on X, and $f : X \to \mathbb{R}$ is a fitness (or objective) function.

A point $x \in X$ is a *local optimum* (of f on X) if

 $f(y) \le f(x) \quad \forall \, yRx$

and a global optimum (maximum) if

 $f(y) \le f(x) \quad \forall \ y \in X$

Questions of the "ruggedness" of landscapes (correlation structure), number and height of local optima, sizes of "attraction basins" of local optima with respect to "hill-climbing" algorithms etc. are of great interest for natural landscapes.

In Kauffman's NK models, $X = A^N$ (usually just $X = \{0, 1\}^N$) and K is a tunable neighbourhood size parameter that influences the landscape characteristics, especially its ruggedness (cf. Figure 4).

The model can be seen as a toy model of "epigenetic interactions in chromosomes" — or also a generalisation of the spin glass model.

In Kauffman's model, a *chromosome* is a *N*-vector of *loci* (*genes*, "positions"), each of which has a value from a set of *allelles A* (usually just $A = \{0, 1\}$). A "filled-in" chromosome $\alpha \in A^N$ is called a *genotype*.

The fitness of each gene $i \in \{1, ..., N\}$ in a genotype $\alpha = (a_1, ..., a_N) \in A^N$ depends on the allele a_i and K other alleles $a_1^i, ..., a_K^i$ via some local fitness function $f^i(\alpha) = f^i(a_i; a_1^i, ..., a_K^i)$, usually normalised so that $f^i(\alpha) \in [0, 1]$. The total fitness of a genotype $\alpha \in A^N$ is the normalised sum of its genes' local fitnesses:

$$f(\alpha) = \frac{1}{N} \sum_{i=1}^{N} f^{i}(a_{i}; a_{1}^{i}, \dots, a_{k}^{i}) \in [0, 1].$$



Figure 4: A smooth (a) and a rugged (b) NK fitness landscape.



Figure 5: An NK interaction network with N = 5, K = 2.

Figure 5 illustrates an NK network with five loci and two "epigenetic interactions" per locus.

In Kauffman's versions of the model, the *K* loci affecting locus *i* can either be systematically selected as e.g. $i + 1, ..., i + K \pmod{N}$, or the chromosome can be simply "randomly wired". The f^i are usually determined as randomly generated 2^{K+1} -element "interaction tables".

From the spin glass perspective, e.g. a 1-D Ising model with N spins can be seen as an N2 network where $f^i(S_i; S_{i-1}, S_{i+1}) = \frac{J}{2}(S_{i-1}S_i + S_iS_{i+1})$, and an SK spin glass with coefficients J_{ij} and local fields h_i as an N(N-1) network where

$$f^{i}(S_{i}; \sigma \setminus \{S_{i}\}) = \frac{1}{2} \sum_{\langle ij \rangle} J_{ij} S_{i} S_{j} + h_{i} S_{i}.$$

Basic properties of the NK model, for binary alleles $A = \{0, 1\}$ and varying values of *K*, include the following:

K = 0:

If $f^i(0) \neq f^i(1) \forall i = 1, ..., N$, then there is a unique global optimum, which is easily found by e.g. the obvious 1-locus mutation "hill-climbing" algorithm.

Expected length of the hill-climbing path is N/2. (Half of the alleles are "right" in the beginning, after that one allele gets fixed at each step.)

Neighbouring genotypes α , α' are always highly correlated, as necessarily $|f(\alpha) - f(\alpha')| \le 1/N$.

 $1 \le K < N - 1$:

For K = 1, a global optimum can still be found in polynomial time. For $K \ge 2$, global optimisation is NP-complete. However, for adjacent affecting loci $(i \curvearrowleft i+1, \ldots, i+K)$, the problem can be solved in time $o(2^K N)$ (Weinberger).

K = N - 1:

Neighbouring genotypes are totally uncorrelated.

 \Rightarrow Probability that a given genotype α is a local optimum is equal to the probability that α has the highest rank within its 1-mutant neighbourhood. This probability is equal to 1/(N+1).

 \Rightarrow The expected number of local optima is $2^N/(N+1)$.

The expected number of improvement steps for 1-mutant hill-climbing to hit a local optimum is proportional to $\log_2 N$ (each improvement step typically halves the rank of the genotype within the neighbourhood).

The expected waiting time for finding an improvement step is proportional to N.

7 Random Graphs

7.1 The Erdős-Rényi Model(s)

Two closely related "uniform" random graph models introduced in 1959 by P. Erdős & A. Rényi and E. N. Gilbert.

Consider the family g_n of all (labelled, undirected) graphs on *n* nodes. Denote $N = \binom{n}{2}$; then $|g_n| = 2^N$.

Define the following two probability spaces

[Erdős & Rényi:] $\mathcal{G}(n,M) = \text{all } G \in \mathcal{G}_n$ with exactly $M \leq N$ edges, taken with uniform probability, i.e.

$$\Pr(G_M = H) = \begin{cases} \binom{N}{M}^{-1}, & \text{if } H \text{ has } M \text{ edges} \\ 0; & \text{otherwise.} \end{cases}$$

[Gilbert:] $g(n, p) = \text{all } G \in g_n$, taken so that each edge has occurrence probability $p, 0 \le p \le 1$, independently of the other edges, i.e.

$$\Pr(G_p = H) = p^M \underbrace{(1-p)}_{q}^{N-M}$$
, if *H* has *M* edges.

These spaces are in a precise sense "close" if $M \sim pN$, and are often both referred to (unfairly to Gilbert) as the "Erdős-Rényi random graph model", or alternatively as the $\mathcal{G}(n, M)$ and $\mathcal{G}(n, p)$ random graph models.

Let Ω_n , n = 0, 1, 2, ... be a sequence of probability spaces of *n*-node graphs. Say that *almost every* (*a.e.*) graph in Ω_n has property Q if

 $\Pr(G \in \Omega_n \text{ has } Q) \to 1, \text{ as } n \to \infty.$

Conversely, *almost no* graph in Ω_n has property Q if a.e. graph in Ω_n has property $\neg Q$, i.e.

 $\Pr(G \in \Omega_n \text{ has } Q) \to 0, \text{ as } n \to \infty.$

Theorem 7.1 Let H be a fixed graph and p a constant, $0 . Then a.e. <math>G \in \mathcal{G}(n, p)$ contains an induced copy of H.

Remark: an "induced copy" means here a subset of nodes whose induced subgraphs is isomorphic to H.

Proof. Let k = |H| = number of nodes in H. Then a graph G with $n = |G| \ge k$ nodes can be partitioned into $\lfloor n/k \rfloor$ disjoint sets of k nodes (with some left over). For each of these sets, the probability that it forms an induced copy of H is r > 0. (Precisely, $r = \frac{k!}{|\operatorname{Aut}(H)|} p^{e(H)} q^{\binom{k}{2} - e(H)}$.)

Thus, the probability that none of these sets forms an induced copy of H is

 $(1-r)^{\lfloor n/k \rfloor} \to 0$, as $n \to \infty$.

Let $k, l \in \mathbb{N}$. Say that a graph G = (V, E) has property Q_{kl} if $\forall U, W, |U| \le k, |W| \le l, U \cap W = \emptyset$, *G* contains a node $v \in V$ such that *v* is adjacent to all $u \in U$ and no $w \in W$ (cf. Figure 6).

Lemma 7.2 For every constant p, $0 , and all <math>k, l \in \mathbb{N}$, a.e. $G \in \mathcal{G}(n, p)$ has property Q_{kl} .



Figure 6: Property Q_{kl} .

Proof. For a fixed $U, W, v \in V \setminus (U \cup W)$, the probability that the condition is satisfied is

$$p^{|U|}q^{|W|} \ge p^k q^l$$

The events are independent for different v, so the probability that no appropriate v exists is

$$\left(1 - p^{|U|}q^{|W|}\right)^{n-|U|-|W|} \le \left(1 - p^k q^l\right)^{n-k-l}$$

There are at most $n^{k+l}(U,W)$ -pairs to be considered, so the probability that some pair has no good *v* is bounded by

$$n^{k+l}(\underbrace{1-p^kq^l}_{<1})^{n-k-l} \to 0, \text{ as } n \to \infty.$$

Thus in a.e. $G \in \mathcal{G}(n, p)$ all (U, W)-pairs have some appropriate v. \Box

Corollary 7.3 Let p, $0 , be a constant. Then (i) a.e. <math>G \in \mathcal{G}(n,p)$ has minimum degree $\geq k$, for given constant k (ii) a.e. $G \in \mathcal{G}(n,p)$ has diameter 2 (iii) a.e. $G \in \mathcal{G}(n,p)$ is k-connected for given constant k.

Proof. (i) and (ii) are immediate.

(iii) In a.e. $G \in g(n, p)$, no two nodes u_1 , u_2 can be separated by a cutset of size k-1, because we may choose in Lemma 7.2 $U = u_1, u_2, W = w_1, \ldots, w_{k-1}$ for arbitrary w_1, \ldots, w_{k-1} , and obtain a path $u_1 - v - u_2$ connecting u_1, u_2 and avoiding w_1, \ldots, w_{k-1} . \Box

Corollary 7.4 Let ϕ be any first-order sentence about graphs (i.e. quantification over nodes, relations E(u, v) + identity). Then either $G \models \phi$ or $G \models \neg \phi$ for a.e. $G \in \mathcal{G}(n, p)$.

Proof. Induction on the structure of ϕ , using Lemma 7.2 to eliminate quantifiers. \Box

Thus, all the first-order properties of g(n,p) for fixed p are easily captured. Things are more interesting when the number of nodes discussed and/or the probability p depends on n.

Given graph G, denote:

l for
)

Lemma 7.5 *Given* $n \ge k \ge 2$ *, random* $G \in \mathcal{G}(n, p)$ *:*

$$\Pr(\alpha(G) \ge k) \le {\binom{n}{k}}q^{\binom{k}{2}}.$$

Proof. Probability that given k-set of nodes in G is independent is $q^{\binom{k}{2}}$. Total number of k-sets is $\binom{n}{k}$. \Box

Theorem 7.6 Let $p, 0 and <math>\varepsilon > 0$ be constant. Then for a.e. $G \in \mathcal{G}(n, p)$:

$$\chi(G) > \frac{\ln(1/q)}{2+\varepsilon} \cdot \frac{n}{\ln(n)} = \Omega\left(\frac{n}{\ln(n)}\right) = large!$$

Proof. By Lemma 7.5, for any fixed $n \ge k \ge 2$:

$$Pr(\alpha(G) \ge k) \le {\binom{n}{k}} q^{\binom{k}{2}} \le n^k q^{\binom{k}{2}}$$
$$= q^{k\frac{\ln n}{\ln q} + \frac{1}{2}k(k-1)}$$
$$= q^{\frac{k}{2}[-\frac{2\ln(n)}{\ln(1/q)} + k - 1]}$$
$$\to 0 \text{ for } k \text{ large,}$$

when

$$\frac{k}{2} \left[-\frac{2\ln(n)}{\ln(1/q)} + k - 1 \right] \to \infty \text{ for } k > (2+\varepsilon) \frac{\ln(n)}{\ln(1/q)}.$$

Thus there cannot be a.s. *k* nodes coloured with same colour for large k \Rightarrow More than $\frac{n}{k} = \frac{\ln 1/q}{2+\epsilon} \cdot \frac{n}{\ln n}$ colours needed. \Box

Theorem 7.7 Let p, $0 be constant. Then for a.e. <math>G \in \mathcal{G}(n, p)$:

 $\omega(G) \in \{d, d+1\},\$

where d = d(n, p) is the largest integer such that

$$\binom{n}{d} p^{\binom{d}{2}} \ge \ln n$$

(This implies $d = 2\log_{1/p}(n) + O(\log\log(n).)$

A graph property Q is an isomorphism-closed family of graphs, i.e. if $G \in Q$ (or "G has Q") and $G \approx G'$, then also $G' \in Q$.

A graph property is *monotone* if it is preserved under addition of edges, i.e. if G = (V, E) and G' = (V, E') are graphs such that $E \subseteq E'$ and G has Q, then also G' has Q.

A *threshold function* for the graph property Q is a function $t : \mathbb{N} \to \mathbb{R}$ such that

$$\Pr(G \in \mathcal{G} (n, p(n)) \text{ has } Q) \xrightarrow[n \to \infty]{} \begin{cases} 1, \text{ if } p \succ t \\ 0, \text{ if } p \prec t. \end{cases}$$

Notation:

$$p \succ t \Leftrightarrow \lim_{n \to \infty} \frac{p(n)}{t(n)} = \infty,$$
$$p \prec t \Leftrightarrow \lim_{n \to \infty} \frac{p(n)}{t(n)} = 0,$$
$$p \sim t \Leftrightarrow \lim_{n \to \infty} \frac{p(n)}{t(n)} = 1,$$
$$p \approx t \Leftrightarrow p(n) = \Theta(t(n)).$$

Denote: $P_n^Q(p) = \Pr(G \in \mathcal{G}(n, p) \text{ has } Q).$ Then for monotone $Q: p_1 \leq p_2 \Rightarrow P_n^Q(p_1) \leq P_n^Q(p_2) \forall n.$



Figure 7: $P_n^Q(p)$ for (a) small, (b) intermediate and (c) large *n*.

Denote: $p_n^Q(\alpha)$ = the smallest p such that $P_n^Q(p) \ge \alpha$.

In fact $P_n^Q(p)$ is a continuous, strictly increasing function, so really $p_n^Q(\alpha) =$ unique *p* such that $P_n^Q(p) = \alpha$.

Figure 7 illustrates the evolution of the function P_n^Q , and a corresponding threshold function t(n), for a monotone graph property Q from small to large values of n.

Lemma 7.8 A function t(n) is a threshold for monotone graph property Q if and only if

$$t(n) \approx p_n^Q(\alpha)$$

for all $0 < \alpha < 1$.

Proof.

" \Rightarrow " Assume t(n) is a threshold for Q. This means that if $p(n)/t(n) \rightarrow \infty$, then

$$P_n^Q(p(n)) \to 1 \qquad (*)$$

and if $p(n)/t(n) \rightarrow 0$, then

$$P_n^Q(p(n)) \to 0 \qquad (**)$$

Suppose then that

$$t(n) \not\approx p_n^Q(\alpha)$$

for some $0 < \alpha < 1$. This means that either there is a sequence n_1, n_2, \ldots such that

$$p_{n_k}^Q(\alpha)/t(n_k) \to \infty$$

contradicting (*), or there is a sequence n_1, n_2, \ldots such that

$$p_{n_k}^Q(\alpha)/t(n_k) \to 0$$

contradicting (**).

(Note that by definition, $P_n^Q(p_n^Q(\alpha)) = \alpha!$)

" \Leftarrow " Assume that t(n) is *not* a threshold for Q. Then there is either a sequence n_1, n_2, \ldots such that

$$p(n_k)/t(n_k) \to \infty$$
,

but

$$P_n^Q(p(n_k)) \leq \alpha < 1$$

a.e., or a sequence n_1, n_2, \ldots such that

$$p(n_k)/t(n_k) \to 0,$$

but

$$P_n^Q(p(n_k)) \ge \alpha > 0$$

a.e. In the former case,

$$t(n_k) \prec p(n_k) \leq p_{n_k}^Q(\alpha),$$

and in the latter case

$$t(n_k) \succ p(n_k) \ge p_{n_k}^Q(\alpha).$$

Thus in either case $t(n) \not\approx p_n^q(\alpha)$. \Box

Theorem 7.9 *Every monotone graph property Q has a threshold function.*

Proof. For brevity, denote $p_n^Q(\alpha) = p(\alpha)$. Choose some arbitrary $0 < \alpha < \frac{1}{2}$. The goal is to prove that $p(\alpha) \approx p(1-\alpha)$, thus establishing e.g.

$$t(n) = p\left(\frac{1}{2}\right) = p_n^Q\left(\frac{1}{2}\right)$$

as a threshold function for Q. (Since $p(\alpha) \le p(\frac{1}{2}) \le p(1-\alpha)$.)

Let $m \in \mathbb{N}$ be such that $(1 - \alpha)^m \leq \alpha$. Let $p = p_n(\alpha)$ and consider a sample of *m* independent graphs G_1, \ldots, G_m from $\mathcal{G}(n, p)$. Then the graph $G_1 \cup \cdots \cup G_m \in \mathcal{G}(n,q)$, where $q = 1 - (1-p)^m \leq mp$, and so

$$\Pr(G_1 \cup \cdots \cup G_m \text{ has } Q) \leq \Pr(G \in \mathcal{G}(n, mp_n(\alpha)) \text{ has } Q).$$

On the other hand, since Q is monotone, if any G_i has Q, then so does $G_1 \cup \cdots \cup G_m$. Thus,

$$\Pr(G_1 \cup \cdots \cup G_m \text{ does not have } Q) \le (1 - \Pr(G_i \text{ has } Q))^m \\= (1 - \alpha)^m \le \alpha.$$

Hence,

$$\Pr_n^Q(mp_n(\alpha)) \ge \Pr(G_a \cup \cdots \cup G_m \text{ has } Q) \ge 1 - \alpha,$$

and so

$$p_n(\alpha) \leq p_n(1-\alpha) \leq mp_n(\alpha),$$

i.e. $p(\alpha) \approx p(1-\alpha)$. (Since *m* depends only on α , not on *n*.)

Consider a graph property Q defined as "G has Q" if X(G) > 0, where $X \ge 0$ is a random variable on $\mathcal{G}(n, p)$.

E.g. if X(G) denotes the number of spanning trees of G, then property Q corresponds to connectedness.

A threshold function for property Q is a t(n) such that

- (i) $p(n) \prec t(n) \Rightarrow \text{almost no } G \in G(n, p(n)) \text{ has } Q.$
- (ii) $p(n) \succ t(n) \Rightarrow$ almost all $G \in G(n, p(n))$ have Q.

If X is integral, then condition (i) can be verified by upper bounding E[X]; by Markov's inequality:

$$\Pr(X \ge 1) \le E[X] \qquad (\text{ more generally, for } a > 0$$

$$p(X \ge a) \le E[X]/a).$$

Condition (ii) is trickier, but can be approached by lower-bounding E[X], and upper-bounding Var[X]. (So called "second-moment method".)

Denote
$$\mu = E[X], \sigma^2 = \text{Var}[X] = E[(X - \mu)^2] = E[X^2] - \mu^2.$$

Recall Chebyshev's inequality: for any $\lambda > 0$,

$$\Pr(|X-\mu| \ge \lambda) \le \frac{\sigma^2}{\lambda^2}.$$

Lemma 7.10 If $\mu > 0$ for n large, and $\frac{\sigma^2}{\mu^2} \to 0$ as $n \to \infty$, then X(G) > 0 for a.e. $G \in \mathcal{G}(n, p)$.

Proof. If X(G) = 0, then $|X(G) - \mu| = \mu$. Hence

$$\Pr(X=0) \le \Pr(|X-\mu| \ge \mu) \le \frac{\sigma^2}{\mu^2} \to 0 \text{ as } n \to \infty.$$

Denote the *density* of a graph *G* by $\delta(G) = \frac{e(G)}{|G|}$.

Say that a graph *G* is *balanced* if $\delta(G') \leq \delta(G)$ for all subgraphs *G'* of *G*.

Theorem 7.11 Let *H* be a balanced graph. Then the graph property "*G* has a subgraph isomorphic to *H*" has threshold function $n^{-1/\delta(H)}$.

Proof. Let X(G) =number of H-subgraphs of G. Let k = |H|, l = e(H), so $\delta(H) = l/k$. Let us first bound E[X] from above. Let $G \in \mathcal{G}(n, p)$, where $p = \gamma n^{-1/\delta(H)} = \gamma n^{-k/l}$ for some $\gamma = \gamma_n \to 0$, and denote

 $\mathcal{H} = \{ \text{all copies of } H \text{ on vertex-set of } G \}.$

Then $|\mathcal{H}| = \binom{n}{k}h \le \binom{n}{k}k! \le n^k$. Here *h* is the number of different arrangements of *H* on a set of *k* vertices, $h = k!/|\operatorname{Aut}(H)|$. Thus

$$E[X] = \sum_{H' \subseteq \mathcal{H}} \Pr(H' \in G) = |\mathcal{H}| \cdot p^l \ \leq n^k p^l = n^k (\gamma n^{-k/l})^l = \gamma^l \xrightarrow[\gamma \to 0]{} 0.$$

Thus if

$$p(n) = \gamma_n \cdot n^{-1/\delta(H)} \prec n^{-1/\delta(H)},$$

then $E[X] \xrightarrow[n \to \infty]{} 0.$

By Markov's inequality this means that almost no $G \in \mathcal{G}(n, p)$ contains an *H*-subgraph for large *n*.

For the other part, we need to bound from above

$$\frac{\sigma^2}{\mu^2} = \frac{1}{\mu^2} (E[X^2] - \mu^2).$$

Let us try to compute:

$$E[X^{2}] = \sum_{H',H''\in\mathcal{H}} \Pr(H'\cup H''\subseteq G)$$

=
$$\sum_{H',H''\in\mathcal{H}} p^{e(H')+e(H'')-e(H'\cap H'')}$$

$$\leq \sum_{H',H''\in\mathcal{H}} p^{2l-i\delta(H)},$$

where $i = |H' \cap H''|$. (Note that $\delta(H' \cap H'') \le \delta(H)$.)

Denote then $\mathcal{H}_i^2 = \{(H', H'') \in \mathcal{H}^2 : |H' \cap H''| = i\}$ and compute separately for each *i* the sum

$$A_i = \sum_{\mathcal{H}_i^2} \Pr(H' \cup H'' \subseteq G)$$

Case i = 0:

$$A_{0} = \sum_{\mathcal{H}_{0}^{2}} \Pr(H' \cup H'' \subseteq G)$$

= $\sum_{\mathcal{H}_{0}^{2}} \Pr(H' \subseteq G) \cdot \Pr(H'' \subseteq G)$ H', H'' independent
 $\leq \sum_{\mathcal{H}^{2}} \Pr(H' \subseteq G) \cdot \Pr(H'' \subseteq G)$
= $\left(\sum_{\mathcal{H}} \Pr(H' \subseteq G)\right)^{2}$
= μ^{2}

Case i \geq 1:

$$\begin{split} A_{i} &= \sum_{\mathscr{H}_{i}^{2}} \Pr(H' \cup H'' \subseteq G) \\ &= \sum_{H' \in \mathscr{H}} \sum_{\substack{H'' : \\ |H' \cap H''| = i}} \Pr(H' \cup H'' \subseteq G) \\ &\leq |\mathscr{H}| \cdot \binom{k}{i} \binom{n-k}{k-i} hp^{2l} p^{-il/k} \qquad h = \frac{k!}{|\operatorname{Aut}(H)|} \\ &\leq |\mathscr{H}| \cdot c_{1} n^{k-i} hp^{2l} (\gamma n^{-k/l})^{-il/k} \\ &= \mu \cdot c_{1} n^{k-i} hp^{l} \gamma^{-il/k} n^{i} \\ &= \mu c_{2} \underbrace{\binom{n}{k}} hp^{l} \gamma^{-il/k} \\ &= \mu c_{2} \underbrace{\binom{n}{k}} hp^{l} \gamma^{-il/k} \\ &\leq \mu^{2} \cdot c_{2} \gamma^{-il/k}. \end{split}$$

Thus, denoting $c_3 = kc_2$, we get the estimate

$$\frac{E[X^2]}{\mu^2} = \left(\frac{A_0}{\mu^2} + \frac{\sum_i A_i}{\mu^2}\right) \le 1 + c_3 \gamma^{-l/k}$$

and hence

$$\frac{\sigma^2}{\mu^2} = \frac{E[X^2] - \mu^2}{\mu^2} \le c_3 \gamma^{-l/k} \xrightarrow[\gamma \to \infty]{} 0.$$

Hence, if $p(n) = \gamma_n n^{-k/l}$ for $\gamma_n \to \infty$, then by Lemma 7.10 X(G) > 0 holds for almost every $G \in \mathcal{G}(n, p)$ for large $n.\Box$

Corollary 7.12 For $k \ge 3$, the property of containing a k-cycle has threshold $t(n) = n^{-1}$. (Note: independent of k)

Corollary 7.13 For $k \ge 2$, the property of containing a specific tree structure T on k nodes has threshold function $t(n) = n^{-k/(k-1)}$. \Box

Corollary 7.14 For $k \ge 2$, the property of containing a k-clique ($\approx K_k$) has threshold function $t(n) = n^{-2/(k-1)}$. \Box

Denote $\delta^*(H) = \max{\{\delta(H') | H' \text{ is subgraph of } H\}}$.

Theorem 4.11' The graph property "G has a subgraph isomorphic to H" has threshold function $n^{-1/\delta^*(H)}$.

Threshold functions for global graph properties

Also known as "the phase transition".

The "epochs of evolution": Consider the structure of random graphs $G \in \mathcal{G}(n, p)$, as p = p(n) increases. The following results can be shown (note that np = average node degree):

- 0. If $p \prec n^{-2}$, then a.e. *G* is empty.
- 1. If $n^{-2} \prec p \prec n^{-1}$, then a.e. *G* is a forest (a collection of trees).
 - The threshold for the appearance of any *k*-node tree structure is $p = n^{-k/(k-1)}$.