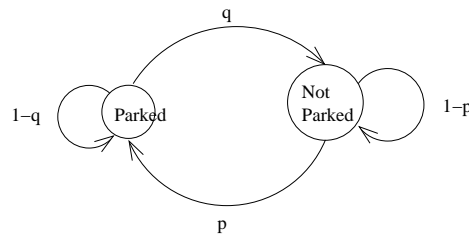


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Solutions

1. A Fire station has one fire truck. When an emergency call comes, the truck goes out, fights fire and then returns.



We use shorthands p and q for the transition probabilities of the HMM illustrated in the figure.

- (a) First, we need to figure out values for p and q .

For each one-hour time slice, the probability that the truck returns is p and the probability that it does not return is $1 - p$. Thus the expected time that the truck is away is

$$\sum_{i=0}^{\infty} (i+1) \cdot p \cdot (1-p)^i = \left[\sum_{i=0}^{\infty} (1-p)^i \right]^2 = \frac{p}{(1-(1-p))^2} = \frac{1}{p}$$

hours. By equating this with three hours, we obtain $p = \frac{1}{3}$.

On the other hand, we assign $q = \frac{1}{12}$, as there is an emergency call once in twelve hours on the average.

- (b) $\mathbf{P}(Parked_{t+1} \mid Parked_t) = \langle \frac{11}{12}, \frac{1}{12} \rangle$ and
 $\mathbf{P}(Parked_{t+1} \mid \neg Parked_t) = \langle \frac{1}{3}, \frac{2}{3} \rangle$.
- (c) Suppose that $\mathbf{P}(Parked_t) = \langle r, 1-r \rangle$ where r is introduced as a parameter probability. Using the transition model, we obtain

$$\begin{aligned} \mathbf{P}(Parked_{t+1}) &= r \cdot \langle 1-q, q \rangle + (1-r) \cdot \langle p, 1-p \rangle \\ &= \langle r \cdot (1-q) + (1-r) \cdot p, qr + (1-r) \cdot (1-p) \rangle \\ &= \langle r \cdot (1-p-q) + p, (1-r) \cdot (1-p-q) + q \rangle. \end{aligned}$$

In the long run, we have $\mathbf{P}(Parked_t) = \mathbf{P}(Parked_{t+1})$ from which we obtain $r = r \cdot (1-p-q) + p$. It follows that $r \cdot (1-1+p+q) = p$ and $r = \frac{p}{p+q}$. By substituting the known values for p and q , we obtain

$$r = \frac{\frac{1}{3}}{\frac{1}{3} + \frac{1}{12}} = \frac{4}{4+1} = \frac{4}{5}.$$

Thus the truck spends on the average $24 \cdot \frac{4}{5}$ hours (i.e. 19 hours and 12 minutes) per day at the fire station in the long run.

- (d) Our task is to derive an exact expression distribution $\mathbf{P}(Parked_t)$ when $\mathbf{P}(Parked_0) = \langle k, 1 - k \rangle$. We may utilize the general transition probabilities from (c) and conclude the following probabilities:

$$\begin{aligned}
P(Parked_0) &= k, \\
P(Parked_1) &= k \cdot (1 - p - q) + p, \\
P(Parked_2) &= (k \cdot (1 - p - q) + p) \cdot (1 - p - q) + p \\
&= k \cdot (1 - p - q)^2 + p \cdot (1 - p - q) + p, \\
&\dots \\
P(Parked_t) &= k \cdot (1 - p - q)^t + p \cdot \sum_{i=0}^{t-1} (1 - p - q)^i.
\end{aligned}$$

The case of $P(\neg Parked_t)$ can be handled by analogy and symmetry. By exchanging the roles of p and q as well as k and $k - 1$, we obtain

$$P(\neg Parked_t) = (1 - k) \cdot (1 - p - q)^t + q \cdot \sum_{i=0}^{t-1} (1 - p - q)^i.$$

Then we conclude that $\mathbf{P}(Parked_t) = \langle P(Parked_t), P(\neg Parked_t) \rangle$.

- (e) First, it is easy to see from the expressions above that whenever $|1 - p - q| < 1$ or (i.e. $0 < p + q < 2$) it holds that

$$\begin{aligned}
\lim_{t \rightarrow \infty} \mathbf{P}(Parked_t) &= \left\langle p \cdot \frac{1}{1 - (1 - p - q)}, q \cdot \frac{1}{1 - (1 - p - q)} \right\rangle \\
&= \left\langle \frac{p}{p + q}, \frac{q}{p + q} \right\rangle \\
&= \langle r, 1 - r \rangle
\end{aligned}$$

for the probability r derived in (c). In particular, if $1 - p - q = 0$, or equivalently $p + q = 1$, then we have $\lim_{t \rightarrow \infty} \mathbf{P}(Parked_t) = \langle p, q \rangle$. Second, if $1 - p - q = 1$, then $p = q = 0$ and

$$\lim_{t \rightarrow \infty} \mathbf{P}(Parked_t) = \langle k, k - 1 \rangle.$$

Third, if $1 - p - q = -1$, then $p + q = 2$, $p = q = 1$, and we obtain $\mathbf{P}(Parked_t) = \langle k, 1 - k \rangle$ for even values of t and $\mathbf{P}(Parked_t) = \langle 1 - k, k \rangle$ for odd values of t . Thus $\mathbf{P}(Parked_t)$ converges only if $k = 1 - k = \frac{1}{2}$ and $\lim_{t \rightarrow \infty} \mathbf{P}(Parked_t) = \langle \frac{1}{2}, \frac{1}{2} \rangle$.